



A second order Poincaré inequality for functionals of general Poisson processes

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joint work with

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1. Fock space representation of Poisson functionals

Setting

η is a **Poisson process** on some measurable space $(\mathbb{X}, \mathcal{X})$ with **intensity measure** λ . This is a random element in the space \mathbf{N} of all integer-valued σ -finite measures on \mathbb{X} , equipped with the usual σ -field with the following two properties:

- The random variables $\eta(B_1), \dots, \eta(B_m)$ are stochastically independent whenever B_1, \dots, B_m are measurable and pairwise disjoint.



$$\mathbb{P}(\eta(B) = k) = \frac{\lambda(B)^k}{k!} \exp[-\lambda(B)], \quad k \in \mathbb{N}_0, B \in \mathcal{X},$$

where $\infty^k e^{-\infty} := 0$ for all $k \in \mathbb{N}_0$.

Definition (Difference operator)

For a measurable function $f : \mathbf{N} \rightarrow \mathbb{R}$ and $x \in \mathbb{X}$ we define a function $D_x f : \mathbf{N} \rightarrow \mathbb{R}$ by

$$D_x f(\mu) := f(\mu + \delta_x) - f(\mu).$$

For $x_1, \dots, x_n \in \mathbb{X}$ we define $D_{x_1, \dots, x_n}^n f : \mathbf{N} \rightarrow \mathbb{R}$ inductively by

$$D_{x_1, \dots, x_n}^n f := D_{x_1}^1 D_{x_2, \dots, x_n}^{n-1} f,$$

where $D^1 := D$ and $D^0 f = f$.

Definition

Let L^2_η denote the space of all random variables $F \in L^2(\mathbb{P})$ such that $F = f(\eta)$ \mathbb{P} -almost surely, for some measurable function (**representative**) $f : \mathbf{N} \rightarrow \mathbb{R}$. In this case we define

$$D_{x_1, \dots, x_n}^n F := (D_{x_1, \dots, x_n}^n f)(\eta), \quad x_1, \dots, x_n \in \mathbb{X}.$$

Theorem (Ito, Y. '88, L. and Penrose '11)

For any $F, G \in L^2_\eta$,

$$\text{Cov}(F, G) = \sum_{n=1}^{\infty} \frac{1}{n!} \int (\mathbb{E} D_{x_1, \dots, x_n}^n F)(\mathbb{E} D_{x_1, \dots, x_n}^n G) \lambda^n(d(x_1, \dots, x_n)).$$

2. The Poincaré inequality

Theorem (Chen '85; Wu '00; L. and Penrose '11)

For any $F \in L^2_\eta$,

$$\text{Var}[F] \leq \mathbb{E} \int (D_x F)^2 \lambda(dx).$$

Equality holds iff F is a linear function of η .

Theorem

For any $F \in L^2_\eta$,

$$\text{Var}[F] \geq \int (\mathbb{E} D_x F)^2 \lambda(dx).$$

3. A perturbation formula

Theorem (Molchanov and Zuyev '00, L. '14)

Let λ be a σ -finite measure and ν be a finite signed measure such that $\lambda + \nu$ is a measure. Let η_λ and $\eta_{\lambda+\nu}$ be Poisson processes with intensity measure λ and $\lambda + \nu$, respectively. Suppose that $f : \mathbf{N} \rightarrow \mathbb{R}$ is measurable and satisfies a suitable integrability assumption. Then

$$\mathbb{E}f(\eta_{\lambda+\nu}) = \mathbb{E}f(\eta_\lambda) + \sum_{n=1}^{\infty} \frac{1}{n!} \int \mathbb{E}D_{x_1, \dots, x_n}^n f(\eta_\lambda) \nu^n(d(x_1, \dots, x_n)).$$

4. A covariance identity

Definition

For $F \in L^2_\eta$ with representative f we define,

$$P_s F := \int \mathbb{E}[f(\eta^{(s)} + \chi) \mid \eta] \Pi_{(1-s)}(d\chi), \quad s \in [0, 1],$$

where $\eta^{(s)}$ is a **s-thinning** of η and $\Pi_{(1-s)}$ is the distribution of a Poisson process with intensity measure $(1 - s)\lambda$.

Theorem

For any $F, G \in L^2_\eta$ such that $DF, DG \in L^2(\mathbb{P} \otimes \lambda)$,

$$\text{Cov}(F, G) = \mathbb{E} \int_0^1 \int (D_x F)(P_t D_x G) dt \lambda(dx).$$

5. The Stein-Malliavin method

Definition

The **Wasserstein distance** between the laws of two random variables X, Y is defined as

$$d_1(X, Y) = \sup_{h \in \text{Lip}(1)} |\mathbb{E}h(X) - \mathbb{E}h(Y)|.$$

Theorem (Peccati, Solé, Taqqu and Utzet '10)

Suppose that $F \in L^2_\eta$ satisfies $\mathbb{E}F = 0$ and $DF \in L^2(\mathbb{P} \otimes \lambda)$. Let N be standard normal. Then

$$\begin{aligned} d_1(F, N) \leq & \mathbb{E} \left| 1 - \int_0^1 \int (P_t D_x F)(D_x F) dt \lambda(dx) \right| \\ & + \mathbb{E} \int_0^1 \int |P_t D_x F| (D_x F)^2 dt \lambda(dx). \end{aligned}$$



Idea of the proof:

- Use Stein's method:

$$d_1(F, N) \leq \sup_{g \in \mathbf{AC}_{1,2}} |\mathbb{E}[g'(F) - Fg(F)]|.$$

- Use the covariance identity

$$\mathbb{E}Fg(F) = \mathbb{E} \int_0^1 (P_t D_x F)(D_x g(F)) dt \lambda(dx).$$

- Use

$$D_x g(F) = g(F + D_x F) - g(F)$$

and the properties of g .

Remark

Schulte '12 and Eichelsbacher and Thäle '13 derived a similar (but more complicated bound) for the **Kolmogorov distance**

$$d_K(F, N) = \sup_{x \in \mathbb{R}} |\mathbb{P}(F \leq x) - \mathbb{P}(N \leq x)|$$

between the laws of F and N .

6. The second order Poincaré inequality

Theorem (L., Peccati and Schulte '14)

Let $F \in L^2_\eta$ be such $DF \in L^2(\mathbb{P} \otimes \lambda)$, $\mathbb{E}F = 0$ and $\text{Var } F = 1$, and let N be standard normal. Then,

$$d_1(F, N) \leq \gamma_1 + \gamma_2 + \gamma_3,$$

where

$$\gamma_1^2 := 4 \int [\mathbb{E}(D_{x_1} F)^2 (D_{x_2} F)^2]^{1/2} [\mathbb{E}(D_{x_1, x_3}^2 F)^2 (D_{x_2, x_3}^2 F)^2]^{1/2} d\lambda^3,$$

$$\gamma_2^2 := \int \mathbb{E}(D_{x_1, x_3}^2 F)^2 (D_{x_2, x_3}^2 F)^2 d\lambda^3,$$

$$\gamma_3 := \int \mathbb{E}|D_x F|^3 \lambda(dx).$$

Theorem (L., Peccati and Schulte '14)

Let $F \in L^2_\eta$ be such that $DF \in L^2(\mathbb{P} \otimes \lambda)$, $\mathbb{E}F = 0$ and $\text{Var } F = 1$, and let N be standard normal. Then,

$$d_K(F, N) \leq \gamma_1 + \gamma_2 + \gamma_3 + \gamma_4 + \gamma_5 + \gamma_6,$$

where

$$\gamma_4 := \frac{1}{2} [\mathbb{E}F^4]^{1/4} \int [\mathbb{E}(D_x F)^4]^{3/4} \lambda(dx),$$

$$\gamma_5^2 := \int \mathbb{E}(D_x F)^4 \lambda(dx),$$

$$\begin{aligned} \gamma_6^2 := & \int 6 [\mathbb{E}(D_{x_1} F)^4]^{1/2} [\mathbb{E}(D_{x_1, x_2}^2 F)^4]^{1/2} \lambda^2(d(x_1, x_2)) \\ & + \int 3 \mathbb{E}(D_{x_1, x_2}^2 F)^4 \lambda^2(d(x_1, x_2)). \end{aligned}$$

Theorem (L., Peccati and Schulte '14)

For $F, G \in \text{dom } D$ with $\mathbb{E}F = \mathbb{E}G = 0$, we have

$$\begin{aligned} & \mathbb{E} \left(\text{Cov}(F, G) - \int \int_0^1 (D_x F)(P_t D_x G) dt \lambda(dx) \right)^2 \\ & \leq 3 \int [\mathbb{E}(D_{x_1, x_3}^2 F)^2 (D_{x_2, x_3}^2 F)^2]^{1/2} [\mathbb{E}(D_{x_1} G)^2 (D_{x_2} G)^2]^{1/2} d\lambda^3 \\ & \quad + \int [\mathbb{E}(D_{x_1} F)^2 (D_{x_2} F)^2]^{1/2} [\mathbb{E}(D_{x_1, x_3}^2 G)^2 (D_{x_2, x_3}^2 G)^2]^{1/2} d\lambda^3 \\ & \quad + \int [\mathbb{E}(D_{x_1, x_3}^2 F)^2 (D_{x_2, x_3}^2 F)^2]^{1/2} [\mathbb{E}(D_{x_1, x_3}^2 G)^2 (D_{x_2, x_3}^2 G)^2]^{1/2} d\lambda^3. \end{aligned}$$

7. Additive functionals of the Boolean model

Setting

ξ is an **independent marking** of a stationary Poisson process η with intensity $\gamma > 0$. The marks come from the space of all non-empty convex bodies in \mathbb{R}^d and have distribution \mathbb{Q} . The **Boolean model** is the random closed set given by

$$Z := \bigcup_{(x,K) \in \xi} (K + x).$$

We consider a measurable, **additive** and **locally bounded** functional φ on the **convex ring**.

Assumption

We assume that

$$\int \bar{V}(K)^3 \mathbb{Q}(dK) < \infty$$

where $\bar{V} := V_0 + \dots + V_d$ is the sum of intrinsic volumes (**Wills functional**).

Theorem (Hug, L. and Schulte '15+)

Let $W \subset \mathbb{R}^d$ be a convex body and assume that $\sigma_W^2 := \text{Var}[F_W] > 0$, where $F_W := \varphi(Z \cap W)$. Let $\hat{F}_W := \sigma_W^{-1}(F_W - \mathbb{E}[F_W])$. Then

$$d_1(\hat{F}_W, N) \leq c_1 \sigma_W^{-2} \bar{V}(W)^{1/2} + c_2 \sigma_W^{-3} \bar{V}(W),$$

where c_1, c_2 do not depend on W .

Example

In the case $\varphi := V_d$ we have

$$d_1(\hat{F}_W, N) \leq [3(\gamma\gamma_2)^{3/2}c_W + \gamma\gamma_3^{1/2}(c_W)^{3/2}] V_d(W)^{-1/2},$$

where $\gamma_i := \int V_d(K)^i \mathbb{Q}(dK)$ and

$$c_W := V_d(W)(1-p)^{-2} \left[\int V_d(W \cap (W+x)) (e^{\gamma\beta_d(x)} - 1) dx \right]^{-1}.$$

Here $p = 1 - \exp[-\gamma\gamma_1]$ is the **volume fraction** of Z and $\beta_d(x) := \int \lambda_d(K \cap (K+x)) \mathbb{Q}(dK)$.

8. Clusters of the Gilbert graph

Setting

ξ is an **independent marking** of a stationary Poisson process η with intensity $\gamma > 0$. The marks are non-negative random variables (radii). The **Gilbert graph** has vertex set η and an edge between two different points $x, y \in \eta$ if the balls centred at x and y overlap. A **k -cluster** is a connected component with k points.

Theorem

Let $k \geq 1$ and F_W be the number of k -clusters within a compact observation window W . Let $\hat{F}_W := \text{Var}[F_W]^{-1}(F_W - \mathbb{E}[F_W])$. Then, under a suitable (polynomial) moment assumption on the radius distribution,

$$\max\{d_K(\hat{F}_W, N), d_1(\hat{F}_W, N)\} \leq c\lambda_d(W)^{-1/2},$$

where $c > 0$ does not depend on W .

Remark

Penrose '03 proved the (multivariate) CLT for the Gilbert graph with deterministic radii.

9. Cluster counting in the random connection model (Work in progress)

Setting

Let η be a stationary Poisson process with intensity $\gamma > 0$ and let $\varphi : \mathbb{R}^d \rightarrow [0, 1]$ be a measurable and symmetric **connection function**. Given η , connect any two points $x, y \in \eta$ with probability $\varphi(x - y)$ independently of all other pairs. This gives the **random connection model** (η, χ) , where χ is the point process of edges.

Theorem (van de Brug and Meester '04)

Suppose that $\varphi(x) = \tilde{\varphi}(|x|)$ for a decreasing function $\tilde{\varphi}$ satisfying $\int r^{d-1} \tilde{\varphi}(r) dr < \infty$. Let F_W be the number of isolated points within a compact observation window W . Let $\hat{F}_W := \text{Var}[F_W]^{-1/2}(F_W - \mathbb{E}[F_W])$. Then $\hat{F}_W \xrightarrow{d} N$ as $W \uparrow \mathbb{R}^d$. If φ has bounded support, then the same holds for the number of k -clusters.

Goal

The number F_W of k -clusters within a compact observation window W satisfies

$$\max\{d_1(\hat{F}_W, N), d_K(\hat{F}_W, N)\} \leq c\lambda_d(W)^{-1/2}.$$

This should hold under minimal assumptions on φ and also for other suitable localizing (stabilizing) functionals F_W .

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