Differential equations and empirical likelihood

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General formulation

Observed data:

$$\mathbf{Y}_i = \mathbf{X}(t_i) + \varepsilon_i, \qquad i = 1, \dots, n;$$
 (1)

where $t_i \in [T_0, T_1]$.

Parametric ODE model:

$$\frac{d}{dt}\mathbf{X}(t) = G(\mathbf{X}(t), t; \theta)$$
(2)

subject to an initial condition $\mathbf{X}(T_0) = \mathbf{X}_0$.

• If $G(\mathbf{x}, t; \theta) \equiv G(\mathbf{x}, \theta)$, then the system is *autonomous*.

Estimating θ by regression

- Treat $\mathbf{X}(t)$ as a function of (θ, X_0) .
- If X₀ is known, the trajectory for X(t) = X(t, θ) can be determined for all t ∈ [T₀, T₁], up to a given level of precision by numerically solving (2), for example by Runge-Kutta method.
- One obvious approach is to estimate θ by solving the least squares problem:

$$\widehat{\theta}^{LS} = \arg\min_{\theta \in \Theta} \sum_{i=1}^{n} w_i \|\mathbf{Y}_i - \mathbf{X}(t_i; \theta)\|^2$$
(3)

based on some known weights $\{w_i\}_{i=1}^n$.

Computational bottleneck

- Typically, (3) is solved by iteratively solving a sequence of linearized regression problem (e.g., Levenberg-Marquardt method). This involves, at each step of the iterations,
 - 1. Solve for $\mathbf{X}(t; \theta)$ on $[T_0, T_1]$ given the current value of θ .
 - 2. Solve for $\frac{d}{dt}\mathbf{X}(t;\theta)$ on the same interval, by solving the corresponding ODE.
- Both these steps need knowledge of X₀, and evaluation of the trajectory on a fine grid.

Alternative approaches : Two-stage estimation

- The first stage involves estimating X(t) and X'(t) separately using local polynomial method.
- The second stage involves nonlinear regression based on the following description:

$$\widehat{\mathbf{X}}'(t_i) = G(\widehat{\mathbf{X}}(t_i), t_i; \theta) + \mathbf{e}_i, \qquad i = 1, \dots, n.$$

- The method has been analyzed in detail by Chen and Wu (2008a, 2008b).
- ► Choice of weights in the nonlinear regression can impact the quality of estimation. Brunel (2008) (who used splines rather than local polynomial regression for smoothing X(t)), established √n-consistency of θ̂ under a careful choice of weights that vanish at the boundaries.

Alternative approaches : Parameter cascading

- A penalized least squares method whereby the trajectory X is represented in terms of splines and the ODE itself acts as a penalty on the trajectory. The parameters of the ODE are fitted using a parameter cascading method (Ramsay, Hooker, Campbell and Cao, 2007).
- Specifically, after expressing $\widehat{X}(t) = \sum_{j=1}^{M} \beta_j \Phi_j(t)$, they minimize

$$\sum_{i=1}^{n} w_i \|\mathbf{Y}_i - \widehat{\mathbf{X}}(t_i)\|^2 + \lambda \int \|\widehat{\mathbf{X}}'(t) - G(\widehat{\mathbf{X}}(t), t; \theta)\|^2 dt$$

for some $\lambda > 0$.

 Qi and Zhao (2010) provided theoretical analysis of the parameter cascading method.

Alternative approaches : Linear functionals

- Hall and Ma (2014) proposed another smoothing-based estimator by making use of a class of linear test functionals.
- Specifically, if ψ_j, j = 1,..., J are C¹ functions supported on [c, 1 − c] for some c ∈ (0, 1/2), and satisfying ψ_j(c) = ψ_j(1 − c) = 0, then estimate θ by minimizing

$$\sum_{j=1}^{J} W_{j} \| \int_{c}^{1-c} \left(\psi_{j}(t) \widehat{\mathbf{X}}(t) + \psi_{j}(t) G(\widehat{\mathbf{X}}(t), t; \theta) \right) dt \|^{2}$$

where $\widehat{\mathbf{X}}(t)$ is obtained by local polynomial regression.

► They show that the estimator is √n-consistent under appropriate range of the bandwidth and appropriately chosen {\u03c8\u03c9_{j=1}^J.

Empirical likelihood framework

- The main idea is to combine the nonparametric smoothing with the parametric regression of $\widehat{\mathbf{X}}'_{\mathbf{h}}(t)$ on $\widehat{\mathbf{X}}_{\mathbf{h}}(t)$, so that the bandwidths of smoothers can be chosen simultaneously.
- EL step: Use the following normal equations as the "core estimating equations":

$$\sum_{i=1}^{n} w_{i} \frac{\partial}{\partial \theta} \|\widehat{\mathbf{X}}_{\mathbf{h}}^{\prime}(t_{i}) - G(\widehat{\mathbf{X}}_{\mathbf{h}}(t_{i}), t_{i}; \theta)\|^{2} = 0$$
(4)

where $\mathbf{w} := (w_i)_{i=1}^n \in S_n$ (*n*-dimensional simplex).

 Need additional estimating equations to control the bandwidths h and improve stability.

Parametric integral constraint

We impose the restriction

$$\sum_{i=1}^{n} w_i\left(\widehat{\mathbf{X}}''(t_i) - \frac{\partial}{\partial t} G(\widehat{\mathbf{X}}(t), t; \theta) |_{t=t_i}\right) = 0.$$
 (5)

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Maximize Πⁿ_{i=1} w_i subject to (4) and (5), given θ and h, to obtain

$$\ell_P(heta, \mathbf{h}) := \sum_{i=1}^n \log \hat{w}_i(heta, \mathbf{h}).$$

Nonparametric integral constraints

• Given a set of weights $\boldsymbol{\nu} := (\nu_i)_{i=1}^n \in \mathcal{S}_n$, we require:

$$\sum_{i=1}^{n} \nu_i (\mathbf{Y}_i - \widehat{\mathbf{X}}(t_i)) = 0.$$
 (6)

 Based on bias-variance trade-off used to select the optimal global bandwidth, we also impose:

$$\sum_{i=1}^{n} \nu_i \left((Y_{j,i} - \widehat{X}_{j,h_j}(t_i))^2 - c(K_j) n h_j^5 (\widehat{X}_{j,h_j}''(t_i))^2 \right) = 0, \quad (7)$$

for j = 1, *Idots*, *d*, where $c(K_j)$ is a constant determined by the kernel K_j used in obtaining the smoother \hat{X}_{j,h_j} .

• Maximize $\prod_{i=1}^{n} \nu_i$ subject to (4) and (5), given **h**, to obtain

$$\ell_{NP}(\mathbf{h}) := \sum_{i=1}^n \log \hat{
u}_i(\mathbf{h})$$

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Final penalized loss function

We minimize the following loss function to obtain the final estimates of θ and h.

$$L(\theta, \mathbf{h}) = -\ell_P(\theta, \mathbf{h}) - \ell_{NP}(\mathbf{h}) + \eta \sum_{j=1}^d \operatorname{trace}(\mathbf{S}_{j, h_j}^{(1)}) \quad (8)$$

where $\eta \ge 0$ is a specified constant, and $\mathbf{S}_{j,h_j}^{(k)}$ denotes the smoother matrix used to obtain $\widehat{X}_{j,h_j}^{(k)}(t)$ for k = 0, 1, 2.

We choose either η = 0 (no trace-based penalty), or η = 2 (motivated by AIC).

Simulation study

We use the following *Lotka-Volterra* system as a testbed for our procedure. The same system has been studied by Brunel (2008) and Ramsay et al. (2007). $\mathbf{X}(t) = (x(t), y(t))^T$, where

$$\begin{array}{lll} x'(t) &=& \theta_1 x(t) - \theta_3 x(t) y(t) \\ y'(t) &=& -\theta_2 y(t) + \theta_4 x(t) y(t). \end{array}$$

Observational errors:

$$\boldsymbol{\varepsilon}_i = (\varepsilon_i^{\mathsf{x}}, \varepsilon_i^{\mathsf{y}})^{\mathsf{T}} \stackrel{i.i.d.}{\sim} \mathsf{N}(0, \sigma^2 I_2).$$

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Detailed settings

- We compare with the results obtained by Brunel (2008). So, we choose an interval $[T_0, T_1]$ such that $T_1 T_0 = 20$. Also, we choose $(x(T_0), y(T_0)) = (1, 3)$, $(\theta_1, \theta_2, \theta_3, \theta_4) = (1, 1.5, 1.5, 2)$ and $\sigma = 0.2$.
- ▶ We report the results when we have n = 200 equally spaced time points in the interval.
- ▶ The proposed procedure requires a decent initial estimator for fast convergence. We use a two-step procedure where the first step involves estimation of (x(t), y(t)) and their derivatives by local quadratic regression. However, we use cross-validated bandwidths (h_1^{opt}, h_2^{opt}) for obtaining $(\hat{x}(t), \hat{y}(t))$, and $(h_1^{opt}/2, h_2^{opt}/2)$ for obtaining $(\hat{x}'(t), \hat{y}'(t))$.

Parameter	Criterion	Brunel	Initial	Proposed $(\eta = 0)$	$Proposed(\eta=2)$
θ_1	Bias	08	0550	0693	0345
	(SD)	(.07)	(.0677)	(.0806)	(.0709)
θ_2	Bias	-0.10	1422	0950	0716
	(SD)	(.10)	(.1358)	(.1266)	(.1110)
θ_3	Bias	13	1047	1117	0475
	(SD)	(.09)	(.0718)	(.1040)	(.0813)
$ heta_4$	Bias	17	1845	1410	0831
	(SD)	(.12)	(.1130)	(.1524)	(.1073)
σ_x^2	Bias		0135	0204	0185
	(SD)		(.0035)	(.0028)	(.0051)
σ_v^2	Bias		0132	-0.0215	0193
5	(SD)		(.0036)	(.0028)	(.0048)
\hat{h}_{x}	Mean		.0106	.0071	.0079
\hat{h}_y	Mean		.0106	.0068	.0076

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Summary and future work

- Empirical likelihood method provides a flexible approach for estimating ODEs without having to evaluate the trajectories explicitly.
- Appropriate penalization scheme is needed to adjust for bandwidth, which can also be incorporated by combining the parametric and nonparametric aspects.

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- The method enjoys a certain robustness to noise characteristics.
- Asymptotic theory needs to developed.

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