

On the convergence of the spectral empirical process of Wigner matrices

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- ① **Introduction**
- ② **Asymptotic covariances for the ESD (Gaussian U.E. case)**
- ③ **Functional CLT for a spectral empirical process**

1. Introduction: Wigner matrices

The real orthogonal ensemble (O.E.)

- ❶ $W_n = (x_{ij})$, size $n \times n$ symmetric matrix with **independent, zero-mean** entries
- ❷ Moments of second order
 - $\mathbb{E}|x_{ii}|^2 = \sigma^2 > 0$
 - for $i < j$, $\mathbb{E}|x_{ij}|^2 = 1$

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The complex unitary ensemble (U.E.)

- ❶ $W_n = (x_{ij})$, size $n \times n$ Hermitian matrix with independent, zero-mean entries
- ❷ Second moments:
 - $\mathbb{E}|x_{ii}|^2 = \sigma^2 > 0$
 - for $i < j$, $\mathbb{E}|x_{ij}|^2 = 1$ and $\mathbb{E}x_{ij}^2 = 0$

Wigner's semi-circle law

Empirical spectral distribution F_n

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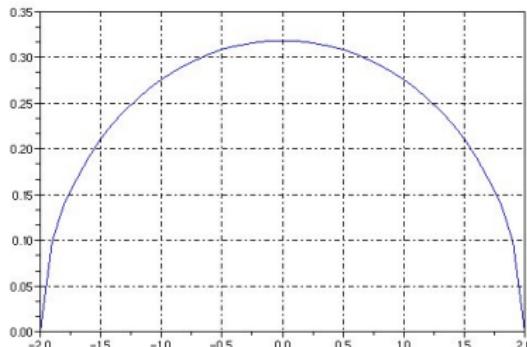
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Theorem

$E F_n$ converges to the semi-circle law, $F(dx) = \frac{1}{2\pi} \sqrt{4 - x^2} dx, \quad x \in [-2, 2]$



1. Introduction: *Convergence rate, CLT*

Other types of LLN

Conditions are known for $F_n \xrightarrow{a.s.} F$ or for $F_n \xrightarrow{P} F$: Review by Bai (1999)

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How about a CLT ?

- First $\sqrt{n}[F_n(x) - F(x)]$ does not converge in distribution
- \sqrt{n} is an uncorrect norming in the “buddle of the spectrum”

$$Var[F_n(x) - F_n(x')] = \frac{2}{\pi^2} \frac{\log n}{n^2} + O(n^{-2}), \quad -2 < x < x' < 2 \quad (0.1)$$

Delyon and Yao, 2004, Gaussian U.E. case

2. Asymptotic covariances for the ESD: Gaussian U.E.

More details:

- joint distribution of the eigenvalues x_1, \dots, x_n of $W_n/\sqrt{2}$:

$$P(x_1, \dots, x_n) = \exp \left\{ - \sum_i x_i^2 + 2 \sum_{i < j} \log |x_j - x_i| + C \right\}$$

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- marginal distribution for two eigenvalues x, y of $W_n/\sqrt{2}$

$$s_n(x, y) = \frac{1}{n(n-1)} \left(K_n(x, x)K_n(y, y) - K_n(x, y)^2 \right),$$

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where

$$K_n(x, y) = \sum_{j=0}^{n-1} \varphi_j(x)\varphi_j(y) = \frac{\sqrt{n}}{\sqrt{2}(x-y)} (\varphi_n(x)\varphi_{n-1}(y) - \varphi_n(y)\varphi_{n-1}(x)).$$

$$\varphi_n(x) = (2^n n! \sqrt{\pi})^{-1/2} e^{-x^2/2} H_n(x), \quad n \geq 0$$

$H_n(x)$ = n -th Hermite polynomial.

Recall

$$\varphi_n(x) = (2^n n! \sqrt{\pi})^{-1/2} e^{-x^2/2} H_n(x), \quad n \geq 0$$

The Plancherel-Rotach formula

Let $\varepsilon > 0$ be fixed. For $\varepsilon \leq \phi \leq \pi - \varepsilon$, let $x = \sqrt{2n+1} \cos \phi$, then

$$(2n)^{1/4} \varphi_n(x) = \left(\frac{2}{\pi \sin \phi} \right)^{1/2} \sin \left[\frac{2n+1}{4} (\sin 2\phi - 2\phi) + \frac{3\pi}{4} \right] + O(n^{-1}),$$

where the error term $O(n^{-1})$ is uniform on the interval $[\varepsilon, \pi - \varepsilon]$.

Convergence rate the expected E.S.D. for $\mathbb{E}F_n$

Proposition

Let $\eta > 0$ be fixed. We have

- i) The density of a single eigenvalue, say $\lambda_{n,1}$ of $W_n/\sqrt{2n}$ has the expansion

$$p_n(u) = \frac{1}{2\pi} \sqrt{4 - u^2} + O(n^{-1}),$$

where the term $O(n^{-1})$ is uniform for all $|u| \leq 2 - \eta$.

ii)

$$\sup_{\overline{uv} \subset [-2+\eta, 2-\eta]} |\mathbb{E}F_n(\overline{uv}) - F(\overline{uv})| = O(n^{-1}).$$

Variance, covariances asymptotiques

- In the so-called “bulk” of the spectrum: $\overline{uv} = [u, v] \subset (-2, 2)$

$$F_n(\overline{uv}) = \frac{1}{n} \sum_{i=1}^n \mathbf{1}_{u \leq \lambda_i \leq v},$$

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- Let $\overline{ab} \subset (-2, 2)$, $\overline{cd} \subset (-2, 2)$

$$\lim_{n \rightarrow \infty} \frac{n^2 \pi^2}{\log n} \operatorname{cov} \left(\frac{1}{2} F_n(\overline{ab}), \frac{1}{2} F_n(\overline{cd}) \right) = \begin{cases} 1, & \text{if } a = c, \ b = d, \\ \frac{1}{2}, & \text{if } a = c < b < d, \\ \frac{1}{2}, & \text{if } c < a < b = d, \\ -\frac{1}{2}, & \text{if } a < c < d \\ 0, & \text{otherwise} \end{cases}$$

Still an open conjecture:

$$\frac{n}{\sqrt{\log n}} [F_n(\overline{ab}) - F(\overline{ab})] \xrightarrow{w} \mathcal{N}(0, *), \quad (a, b) \subset (-2, 2) \quad ??$$

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A related result: Wieand (2001), Diaconis & Evans (2001)

- M_n Haar-distributed form the unitary group
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- M_n Haar-distributed form the unitary group
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Theorem. As $n \rightarrow \infty$, the finite-dimensional distributions of the processes

$$\frac{n\pi}{\sqrt{\log n}} [F_n(\alpha, \beta) - \mathbb{E}F(\alpha, \beta)]$$

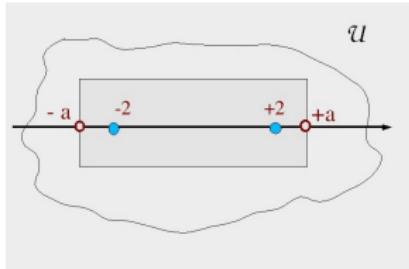
converge to those of a centered Gaussian process $\{Z_{\alpha\beta} : 0 \leq \alpha < \beta < 2\pi\}$ with the covariance structure

$$\text{cov}(Z_{\alpha\beta}, Z_{\alpha'\beta'}) = \begin{cases} 1, & \text{if } \alpha = \alpha' \text{ and } \beta = \beta', \\ \frac{1}{2}, & \text{if } \alpha = \alpha' \text{ and } \beta \neq \beta', \\ \frac{1}{2}, & \text{if } \alpha \neq \alpha' \text{ and } \beta = \beta', \\ -\frac{1}{2}, & \text{if } \beta = \alpha', \\ 0, & \text{otherwise} \end{cases}$$

3. Functional CLT for a spectral empirical process

Bai & Yao (2005)

- Fix an open set $\mathcal{U} \supset [-2, 2]$

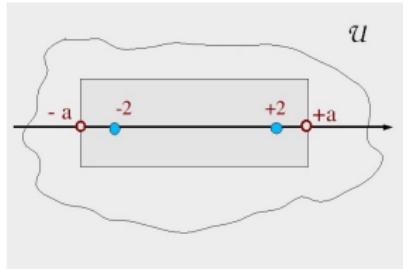


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3. Functional CLT for a spectral empirical process

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\mathcal{A} = analytic functions defined on \mathcal{U}

- Consider integrals

$$G_n(f) := n \int_{-\infty}^{\infty} f(x)[F_n - F](dx), \quad f \in \mathcal{A}.$$

3.a) Main theorem

Main result

- Under suitable conditions, the empirical process

$$G_n := \{ G_n(f), \quad f \in \mathcal{A} \}$$

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- In particular, for any f_1, \dots, f_p of \mathcal{A} ,

$$[G_n(f_1), \dots, G_n(f_p)] \xrightarrow{w} \mathcal{N}_p(0, *).$$

Moment conditions

- [M1] (homogeneity of 4-th moments): $M \equiv \mathbb{E}|x_{ij}|^4$ for $i \neq j$;
- [M2] (uniform tails): for any $\eta > 0$, as $n \rightarrow \infty$,

$$\frac{1}{\eta^4 n^2} \sum_{i,j} \mathbb{E} \left[|x_{ij}|^4 \mathbf{1}_{\{|x_{ij}| \geq \eta \sqrt{n}\}} \right] = o(1).$$

Expansion on Tchebychev polynomials

- $\{T_k\}$ be the family of Tchebychev polynomials
- Orthogonal family:

$$\frac{1}{\pi} \int_{-1}^1 T_i(t) T_j(t) \frac{1}{\sqrt{1-t^2}} dt = \begin{cases} 0, & \text{if } i \neq j \\ 1, & \text{if } i = j = 0 \\ \frac{1}{2}, & \text{if } i = j > 0 \end{cases}$$

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- Expansion on this basis: define

$$\tau_\ell(f) = \frac{1}{\pi} \int_{-1}^1 f(2t) T_\ell(t) \frac{1}{\sqrt{1-t^2}} dt .$$

Theorem

Under the conditions [M1]-[M2],

$$G_n = \{G_n(f), f \in \mathcal{A}\} \xrightarrow{\text{weakly}} G := \{G(f) : f \in \mathcal{A}\}$$

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$$\mathbb{E}[G(f)] = \frac{\kappa - 1}{4} \{f(2) + f(-2)\} - \frac{\kappa - 1}{2} \tau_0(f) + (\sigma^2 - \kappa) \tau_2(f) + \beta \tau_4(f) ,$$

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$$\begin{aligned} \text{cov}(G(f), G(g)) &= (\sigma^2 - \kappa) \tau_1(f) \tau_1(g) + 2\beta \tau_2(f) \tau_2(g) + \kappa \sum_{\ell=1}^{\infty} \ell \tau_{\ell}(f) \tau_{\ell}(g) \\ &= \frac{1}{4\pi^2} \int_{-2}^2 \int_{-2}^2 f'(t) g'(s) V(t, s) dt ds, \end{aligned}$$

where $\kappa = 1$ or 2 , $\beta = \text{var}(|x_{12}|^2) - \kappa$, and

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where $\kappa = 1$ or 2 , $\beta = \text{var}(|x_{12}|^2) - \kappa$, and

$$\begin{aligned} V(t, s) &= (\sigma^2 - \kappa + \frac{1}{2}\beta ts)(4 - t^2)^{1/2}(4 - s^2)^{1/2} \\ &\quad + \kappa \log \left(\frac{4 - ts + (4 - t^2)^{1/2}(4 - s^2)^{1/2}}{4 - ts - (4 - t^2)^{1/2}(4 - s^2)^{1/2}} \right). \end{aligned}$$

3.b). Application: CLT for Linear spectral statistics

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A linear spectral statistic

- f : a polynomial

$$\frac{1}{n}[f(\lambda_{n,1}) + \cdots + f(\lambda_{n,n})] = F_n(f) = \frac{1}{n} \operatorname{tr} f(M_n) .$$

CLT for Tchebychev polynomials

- ① Assume the conditions **[M1]-[M2]** hold.
- ② T_1, \dots, T_p : be p first Tchebychev polynomials.
- ③ Denote $e(k) = \{k \text{ is even}\}$.

Then

$$[G_n(T_1), \dots, G_n(T_p)] \xrightarrow{w} \mathcal{N}_p(w, \Sigma)$$

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with

$$w = (m_\ell), \quad m_\ell = \frac{1}{2} \left[(\kappa - 1)e(\ell) + (\sigma^2 - \kappa)\delta_{\ell 2} + \beta\delta_{\ell 4} \right],$$

$$\Sigma = \text{diag}(a_1, \dots, a_p), \quad a_\ell = \left(\frac{1}{2} \right)^2 \left[(\sigma^2 - \kappa)\delta_{\ell 1} + 2\beta\delta_{\ell 2} + \kappa\ell \right].$$

In particular, these statistics are **asymptotically independent**.

CLT for linear spectral statistics of Gaussian ensembles

The Gaussian U.E.

$$m = \begin{pmatrix} 0 & \vdots & 0 \end{pmatrix}, \quad \Sigma = \begin{pmatrix} (\frac{1}{2})^2 & & & \\ & 2(\frac{1}{2})^2 & \cdots & \\ & & \ddots & \\ & & & k(\frac{1}{2})^2 \\ & & & & \ddots \\ & & & & & p(\frac{1}{2})^2 \end{pmatrix}$$

The Gaussian O.E.

$$m_k = \frac{1}{2}e(k) = \frac{1}{2}\{k \text{ is even }\}, \quad \Sigma = \begin{pmatrix} 2\left(\frac{1}{2}\right)^2 & & & \\ & 4\left(\frac{1}{2}\right)^2 & \cdots & \\ & & \ddots & \\ & & & 2k\left(\frac{1}{2}\right)^2 \\ & & & \ddots & \\ & & & & 2p\left(\frac{1}{2}\right)^2 \end{pmatrix}$$

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- ① Results in this case already provided by (Johansson 1998)
- ② Striking fact: the limit distribution is **non central**.

3.c) Related works on CLT, Gaussian fluctuations

- Starting from [Costin and Lebowitz, 1995]

[Johansson, 1998]

- Extended random ensemble

Joint density of the entries of $W \propto \exp[-n \text{tr}\{V(W)\}]$

V = polynomial of even degree

- Covers G.O.E.
- CLT for linear spectral statistics (Tchebychev polynomials)

[Khorunzhy et al., 1996]

- orthogonal ensemble (with general entries)
- CLT for

$$n\{s_n(z_1) - \mathbb{E}s_n(z_1), \dots, s_n(z_q) - \mathbb{E}s_n(z_q)\}$$

where s_n is the Stieltjes transform

$$s_n(z) = \frac{1}{n} \text{tr} \left(\frac{1}{\sqrt{n}} W_n - zI \right)^{-1}.$$

and

$$z_j \in \mathbb{C}, \quad |\Im(z_j)| \geq 2$$

- CLT very close to fidi convergence part of our theorem

[Sinai and Soshnikov, 1998]

- assume $p = p(n) \rightarrow \infty$, $p/\sqrt{n} \rightarrow 0$.
- the entries are symmetric, having moments of any order with a appropriate growth condition
- CLT for

$$tr \left(\frac{1}{\sqrt{n}} W_n \right)^p - \text{"its expectation"}$$

- Also a CLT for $G_n(f) - \mathbb{E}[G_n(f)]$ with f analytic on the disk of radius 2 without providing the asymptotic mean or variance.

3.d) Some idea of the proofs

Stieltjes transform

- the Stieltjes transform $s_H(z)$ of $H : \mathbb{R} \rightarrow \mathbb{R}$ with **bounded variation**

$$s_H(z) = \int_{-\infty}^{\infty} \frac{dH(x)}{x - z}, \quad z \in \mathbb{C}^+ := \{u + iv, v > 0\}.$$

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- Stieltjes transform of F_n

$$s_n(z) = \frac{1}{n} \sum_{k=1}^n \frac{1}{\lambda_{k,n} - z}.$$

- Stieltjes transform of the semi-circle law F

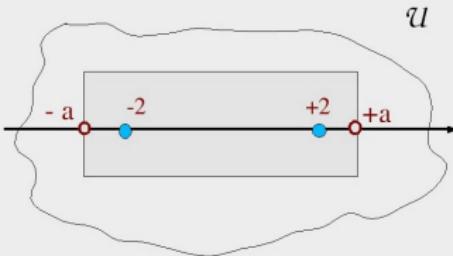
$$s(z) = \frac{1}{2}(z - \sqrt{z^2 - 4}).$$

Integral representation

- Take a contour $\gamma \subset \mathcal{U}$
- for $x \in (-a, a)$,

$$f(x) = \frac{1}{2\pi i} \oint_{\gamma} \frac{f(z)}{z - x} dz .$$

- Then



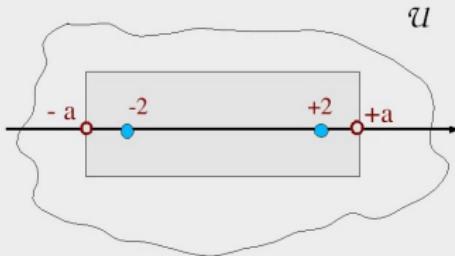
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$$G_n(f) = \int f(x) n[F_n - F](dx)$$



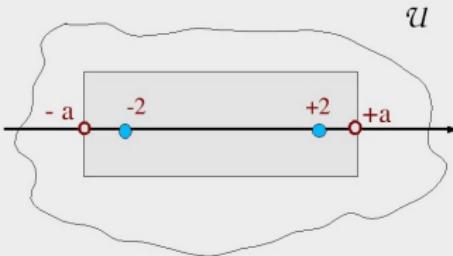
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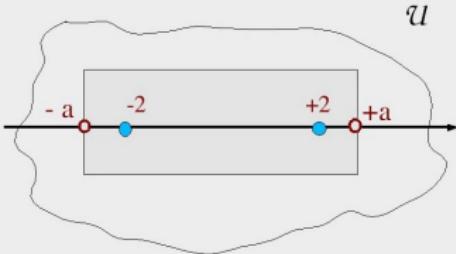
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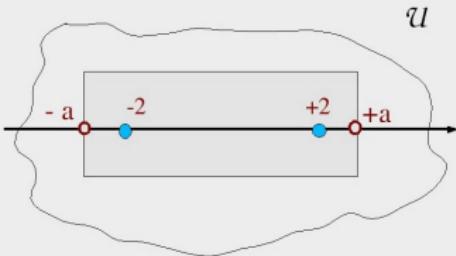
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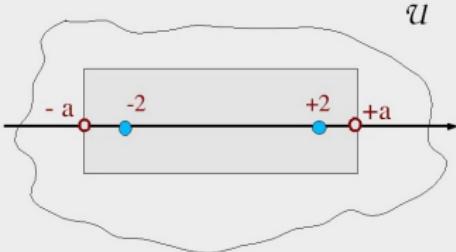
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First reduction: prove the weak convergence of the process

$$\xi_n(z) = n[s_n(z) - s(z)], \quad z \notin [-2, 2]$$

to some Gaussian process.

Proposition

Uniformly on \mathcal{U} , the mean function $\mathbb{E}\xi_n(z)$ tends to

$$\rightarrow [1 + s'(z)]s(z)^3 \left[\sigma^2 - 1 + (\kappa - 1)s'(z)\beta s^2(z) \right]$$

The fidi convergence

Set

$$\begin{aligned}\mathbb{C}_0 &= \{u + iv, |v| \geq v_0\}, \quad v_0 \text{ small,} \\ \zeta_n(z) &= \xi_n(z) - \mathbb{E}\xi_n(z), \quad z \in \mathbb{C}_0\end{aligned}$$

Proposition

Assume Conditions [M1]-[M2] satisfied. For any set of p points $\{z_s, s = 1, \dots, p\}$ of \mathbb{C}_0 , then

$$(\zeta(z_1), \dots, \zeta(z_p)) \Rightarrow \mathcal{N}_p(0, \Gamma),$$

where

$$\begin{aligned}\Gamma(z_j, z_s) &= \frac{\partial^2}{\partial z_j \partial z_s} \left[(\sigma^2 - \kappa)s_j s_s + \frac{1}{2}\beta(s_j s_s)^2 - \kappa \ell n(1 - s_j s_s) \right] \\ &= s'_j s'_s \left[\sigma^2 - \kappa + 2\beta s_j s_s + \frac{\kappa}{(1 - s_j s_s)^2} \right],\end{aligned}$$

with $s_j = s(z_j)$.

Tightness of the sequence of processes $(\zeta_n(z))$

It is enough to establish the following Hölder condition: for some positive constant K

$$\mathbb{E}|\zeta_n(z_1) - \zeta_n(z_2)|^2 \leq K|z_1 - z_2|^2, \quad z_1, z_2 \in \mathbb{C}_0.$$

Conclusions

- ① Provided general weak convergence theorem for the spectral process, indexed by a general class of analytic functions

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- iv) Covers polynomials statistics
- v) Still, I do not know any proof of a result assumed as a folklore by many random matrix people:

$$\frac{n}{\sqrt{\log n}} [F_n(\overline{ab}) - F(\overline{ab})] \xrightarrow{w} \mathcal{N}(0, *), \quad (a, b) \subset (-2, 2) \quad ??$$

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