

# Classification with Ultrahigh-Dimensional Features

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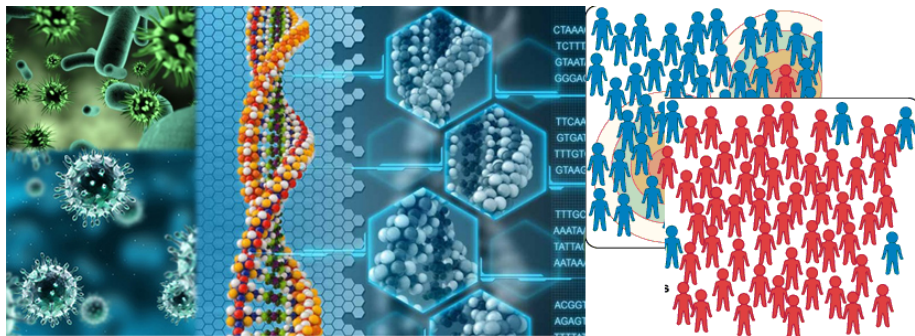
Workshop on Quantitative Methods for Drug Discovery and Development

Singapore, July 2017

- 1 Background
- 2 Motivation and Challenges
- 3 Methods
- 4 Simulation
- 5 A Data Example
- 6 Conclusion

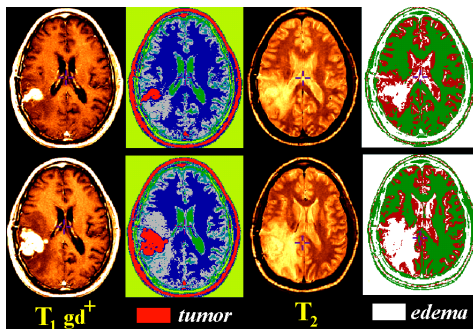
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## From Molecular Diversity to Population Classification



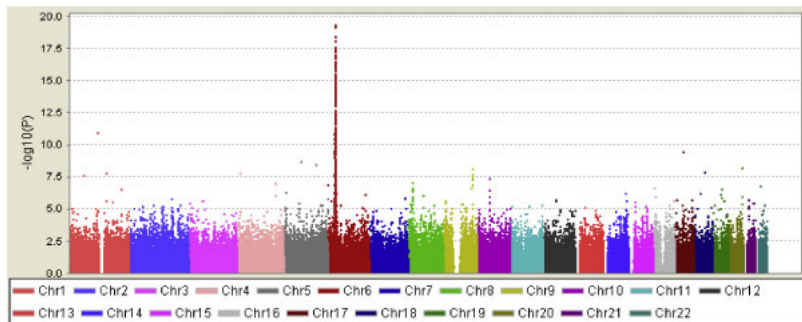
# Some Applications for Ultrahigh-Dimensional Classification

- Distinguish tumor and normal brain tissues



# Applications for Ultrahigh-Dimensional Classification

- Predict disease status based on genomic profiles



# Applications for Ultrahigh-Dimensional Classification

- Classify portfolios based on returns



# Applications for Ultrahigh-Dimensional Classification

- Facial recognition

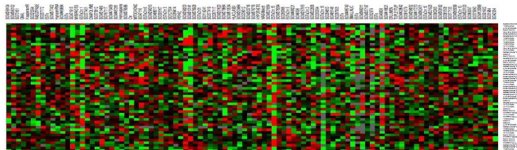




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- 2 Motivation and Challenges**
- 3 Methods
- 4 Simulation
- 5 A Data Example
- 6 Conclusion

# A Motivating Study

- Goal: Classify kidney transplant patients with gene expression profiles
- To distinguish different types of kidney transplant recipients is crucial for
  - balancing immuno-suppression to prevent rejection
  - minimizing drug-introduced toxicities
  - risk prediction for new patients and facilitating personalized transplant policy making
- Dataset: Post-transplant rejection project [[Fletcher et al., 2004](#)] contains 62 kidney transplant patients with 4 post-transplantation rejection types and their genome-wide gene profiles across 12,625 genes



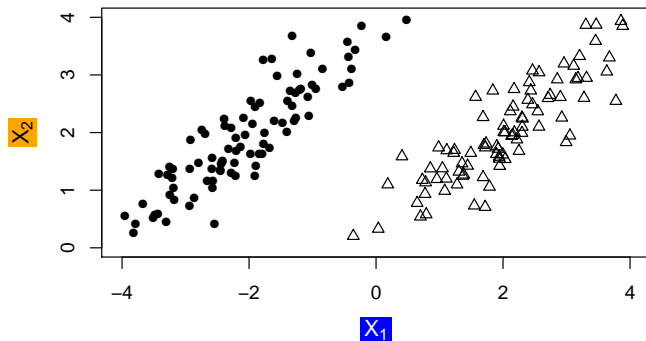
- A common goal is to identify a small number of relevant genetic variants or genes for further investigation
- Screening reduces ultrahigh dimensionality to a handleable scale
- Currently identified genetic variants or genes only explain a small proportion of disease heritability. More relevant features should be included, especially those low-risk susceptibility variants

- Most screening methods focus on detecting marginally strong signals. **Weak signal detection has drawn much attention recently**
- Many classification rules use the independence rule and ignore inter-feature correlations, failing to detect **Marginally Uninformative but Jointly Informative (MUJI)** signals

# Are MUJI Features Famous?



# What is a MUJI Feature?



$X_1$  : Marginally Informative

$X_2$  : Marginally Uninformative but Jointly Informative with  $X_1$

# A Quick Literature Review of High-Dimensional Classification

Methods	Ultrahigh-dimensional screening method other than regularization method?	Consider inter-feature correlation?	Able to detect marginally weak features?
Guo 2010	X	X	X
Witten & Tibshirani 2011	X	X	X
Clemmensen 2011	X	X	X
Fan & Lv 2008	✓	X	X
Xu et al. 2014	X	✓	✓
Cai & Sun 2014	✓	X	X
Cai & Liu 2011	X	✓	✓
Fan, Hong & Tong 2012	X	✓	✓
Mai et al. 2012	X	✓	✓
Gaynanova et al 2016	X	✓	✓

# Proposed Covariance-Insured Screening and Classification Method

- Can handle ultrahigh-dimensional cases
- Utilize inter-feature correlations
- Can detect **MUJI** signals that would be missed by marginal screening methods
- **Can be used to classify features from a much wider class of distributions than Gaussian distribution**
- Computationally efficient
- Recovers all informative features with a large probability
- Has optimal post-screening misclassification error



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- 2 Motivation and Challenges
- 3 Methods**
- 4 Simulation
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- $Y_i$ : a class label taking values in  $\{1, 2\}$
- Conditional on  $Y_i$ ,  $\mathbf{X}_i$  follows a multivariate **sub-Gaussian** distribution with a mean vector  $\boldsymbol{\mu}_k$  and a marginal variance proxy  $\sigma^2$

$$\mathbf{X}_i | \{Y_i = k\} \sim SG(\boldsymbol{\mu}_k, \sigma^2), \quad i = 1, \dots, n \text{ and } k = 1, 2.$$

- Sub-Gaussian distribution contains a wide class of distributions such as Gaussian, binary and all bounded random variables
- Assume that the two classes have a common covariance matrix

$$\boldsymbol{\Sigma} = E[(\mathbf{X}^{(k)} - \boldsymbol{\mu}_k)(\mathbf{X}^{(k)} - \boldsymbol{\mu}_k)']$$

- Most marginal screening processes assume a diagonal  $\boldsymbol{\Sigma}$  and ignore the inter-feature correlations

- Let  $\omega_k = \mathbb{P}(Y_i = k)$ ,  $k = 1, 2$ . For multivariate Gaussian distributed features, consider the pairwise difference

$$\log \mathbb{P}(Y_i = 1 | \mathbf{X}_i = \mathbf{x}) - \log \mathbb{P}(Y_i = 2 | \mathbf{X}_i = \mathbf{x}) = \log \omega_1 - \log \omega_2 \\ - \frac{1}{2} \sum_{j=1}^p (\mu_{1j} + \mu_{2j}) \sum_{j'=1}^p \Omega_{jj'} (\mu_{1j'} - \mu_{2j'}) + \sum_{j=1}^p x_j \left\{ \sum_{j'=1}^p \Omega_{jj'} (\mu_{1j'} - \mu_{2j'}) \right\}$$

- Variable  $j$  is non-informative for distinguishing the two classes if and only if

$$\sum_{j'=1}^p \Omega_{jj'} (\mu_{1j'} - \mu_{2j'}) = 0$$

where  $\Omega_{jj'}$  is the  $(j, j')$ 'th entry of the precision matrix (Xu et al. 2014)

## Definition

Feature  $X_j$  is called a MUJI signal, if and only if it satisfies the following two conditions:

$$\mu_{1j} - \mu_{2j} = 0 \quad (\text{A})$$

$$\sum_{j'=1}^p \Omega_{jj'} (\mu_{1j'} - \mu_{2j'}) \neq 0. \quad (\text{B})$$

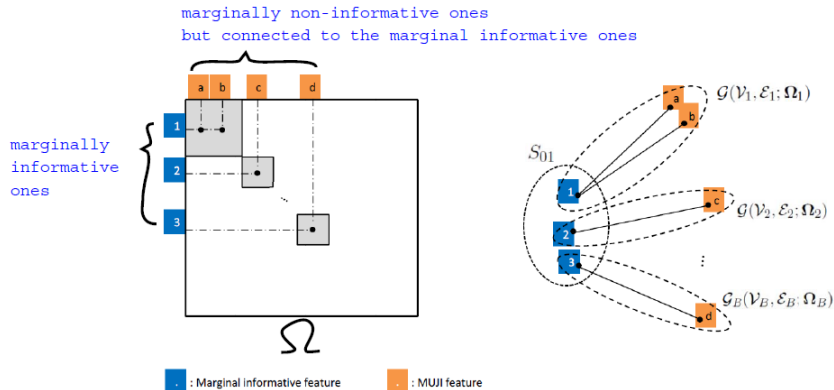
A MUJI feature must be partially correlated with at least one marginally informative feature!

# Challenges in Finding MUJI Features

- Marginal methods do not consider the off-diagonal elements of  $\Omega$
- Estimating  $\Omega$  is infeasible when  $p$  is large

# Our Proposal: Covariance-insured Screening (CIS)

- Identify marginally strong features
- Identify marginally weak/non-informative features that are (partially) correlated with them
- Detecting connected components of marginally informative features in a graph is the Key!



# Detecting Connected Components of a Marginally Informative Feature

- Detect connected components of marginally informative features via thresholding sample covariance matrix  $\tilde{\Sigma}^\alpha \equiv \hat{\Sigma} \mathbb{1}_{\{|\hat{\Sigma}| \geq \alpha\}}$

## Theorem

Let  $\mathcal{G}(\mathcal{V}_j, \mathcal{E}_j; \Omega_j)$  be the connected component of a marginally informative feature  $j$  in  $\mathcal{G}(\mathcal{V}, \mathcal{E}; \Omega)$ , and  $\mathcal{G}(\tilde{\mathcal{V}}_j, \tilde{\mathcal{E}}_j; \tilde{\Sigma}_j^\alpha)$  be the connected component of feature  $j$  in  $\mathcal{G}(\mathcal{V}, \tilde{\mathcal{E}}; \tilde{\Sigma}^\alpha)$ , then for sufficiently large  $n$  and  $\alpha = O(\sqrt{n^{\xi-1}})$ , we have

$$P\left(\mathcal{G}(\mathcal{V}_j, \mathcal{E}_j; \Omega_j) \subseteq \mathcal{G}(\tilde{\mathcal{V}}_j, \tilde{\mathcal{E}}_j; \tilde{\Sigma}_j^\alpha)\right) \geq 1 - C_1 \exp(-C_2 n^\xi),$$

where  $C_1, C_2$  are some positive constants and  $0 < \xi < 1$ .

The goal is to estimate  $\Omega$  or the conditional dependency graph  $\mathcal{E}$

- Given an estimate  $\hat{\Sigma}$  of  $\Sigma$  (which is easy to get!), call variables  $i, j$  conditionally dependent  $\hat{\Sigma}_{ij} \geq \alpha$  (Luo et al., 2014)
- This would give us an estimate of edge set  $\mathcal{E}$
- “From a distance”  $\Omega$  resembles  $\Sigma$ , which makes sense intuitively: the inverse of a block-diagonal matrix is also block-diagonal
- The results hold for sub-Gaussian data



# Covariance-insured Screening and Classification ( $K = 2$ )

## Screening step

- 1 Select marginally informative signals

$$\left| \bar{X}_{\cdot j}^{(1)} - \bar{X}_{\cdot j}^{(2)} \right| = \left| \frac{1}{n_1} \sum_{Y_i=1} X_{ij} - \frac{1}{n_2} \sum_{Y_i=2} X_{ij} \right| > \tau$$

- 2 Construct the thresholded sample covariance matrix

$$\tilde{\Sigma}_{jj'}^{\alpha} = \hat{\Sigma}_{jj'} \mathbb{1}\{|\hat{\Sigma}_{jj'}| \geq \alpha\}, \quad 1 \leq j, j' \leq p$$

- 3 For each selected marginally informative signal, detect its connected component  $\tilde{\Sigma}_k^{\alpha}$  and set the estimated precision matrix

$$\hat{\Omega}^{\alpha, \tau} = \text{diag}((\tilde{\Sigma}_{S_1}^{\alpha})^{-1}, \dots, (\tilde{\Sigma}_{S_{k(\alpha, \tau)}}^{\alpha})^{-1})$$

- 4 Select informative features by

$$S^{\text{CIS}}(\alpha, \tau; \nu_n) = \left\{ j \in \bigcup_{g=1}^{k(\alpha, \tau)} S_g : \left| \sum_{j' \in S_g(j)} (\tilde{\Sigma}_{S_g(j)}^{\alpha})_{jj'}^{-1} (\bar{X}_{\cdot j'}^{(1)} - \bar{X}_{\cdot j'}^{(2)}) \right| \geq \nu_n \right\}$$

## Post-screening classification

5. Classify a new observation  $\mathbf{X}_{\text{new}}$  to class 1 if

$$(\mathbf{X}_{\text{new}} - \hat{\boldsymbol{\mu}}^{\text{CIS}})' \hat{\boldsymbol{\Omega}}^{\text{CIS}} (\hat{\boldsymbol{\mu}}_1^{\text{CIS}} - \hat{\boldsymbol{\mu}}_2^{\text{CIS}}) > 0,$$

where  $\hat{\boldsymbol{\mu}}_1 = \sum_{\{i: Y_i=1\}} \mathbf{X}_i / n_1$ ,  $\hat{\boldsymbol{\mu}}_2 = \sum_{\{i: Y_i=2\}} \mathbf{X}_i / n_2$ ,  $\hat{\boldsymbol{\mu}} = (\hat{\boldsymbol{\mu}}_1 + \hat{\boldsymbol{\mu}}_2) / 2$ ,  $\hat{\boldsymbol{\Omega}} = \hat{\boldsymbol{\Omega}}^{\alpha, \tau}$  and superscript “CIS” restricts features to  $S^{\text{CIS}}(\alpha, \tau; \nu_n)$

- **Importance Score (IS)** evaluates the importance of a feature  $j$ :

$$IS_j = \left| \sum_{j' \in S_{g(j)}} (\tilde{\Sigma}_{S_{g(j)}}^{\alpha})_{jj'}^{-1} (\bar{X}_{\cdot j'}^{(1)} - \bar{X}_{\cdot j'}^{(2)}) \right|$$

- IS incorporates inter-feature correlations. When  $\Omega$  is diagonal, it reduces to the absolute mean differences

- $S_0 = \{j : \sum_{j'=1}^p \Omega_{jj'}(\mu_{1j'} - \mu_{2j'}) \neq 0, j = 1, \dots, p\}$  be the true **informative set**
- $S_{01} = \{j : \mu_{1j} - \mu_{2j} \neq 0, j = 1, \dots, p\}$  is the true **marginally informative set**
- $S_{02} = \{j : \mu_{1j} - \mu_{2j} = 0, \sum_{j'=1}^p \Omega_{jj'}(\mu_{1j'} - \mu_{2j'}) \neq 0, j = 1, \dots, p\}$  is the **MUJI set**
- $S_0 = S_{01} \cup S_{02}$

# A Sparse and Weak Feature Model

- Assume that  $\mu_2$  is only different from  $\mu_1$  in  $\epsilon$  fraction of coordinates, i.e.,  $|S_{01}| = \epsilon p$ .
- Let  $l_1, l_2, \dots, l_p$  be  $p$  samples from  $\text{Bernoulli}(\epsilon)$ , and

$$\mu_{2j} = \mu_{1j} + \delta_j l_j, \quad 1 \leq j \leq p,$$

- Set  $\tau = \min\{|\delta_j| : 1 \leq j \leq p, l_j = 1\} = \sqrt{r \log p}$  for some  $r \in (0, 1)$
- Larger  $r$ , stronger signals
- Set  $\epsilon = p^{-\beta}$  for some  $0 < \beta < 1$ . Larger  $\beta$ , sparser the marginally informative signals
- Assume  $|S_{02}| = p^{1-\gamma}$ , for some  $0 < \gamma < 1$ . Larger  $\gamma$ , sparser the MUJI signals

Let  $S_0 = \{j : \sum_{j'=1}^p \Omega_{jj'}(\mu_{1j'} - \mu_{2j'}) \neq 0, j = 1, \dots, p\}$  be the true informative set (including marginally informative and MUJI features)

## Proposition

With  $\tau = O(\sqrt{r \log p})$ ,  $\alpha = O(\sqrt{n^{\xi-1}})$  and  $\nu_n = O(\sqrt{r \log p})$  and any  $\iota > 0$ ,  $S_{CIS}^{\alpha, \tau}$  satisfies

$$P(|S_0 \cap S_{CIS}^{\alpha, \tau}| \geq (1 - \iota)|S_0|) \rightarrow 1 \text{ as } n \rightarrow \infty$$

## Proposition

For any  $\gamma_n = O(p^\varsigma)$  with  $0 < \varsigma < 1$ ,  $S_{CIS}^{\alpha, \tau}$  satisfies

$$P(|S_{CIS}^{\alpha, \tau} \cap S_0^c| \leq \gamma_n^{-1} |S_0^c|) \rightarrow 1 \text{ as } n \rightarrow \infty$$

- $R_{\text{OPT}}$ : the misclassification rate of the oracle rule assuming that the parameters  $\mu_1$ ,  $\mu_2$  and  $\Sigma$  are known
- $R_{\text{CIS\_PSC}}$ : the misclassification rate for the CIS classification rule  $\mathbf{1}(\hat{\psi}^{\text{CIS}} \geq 0)$  with  $\hat{\psi}^{\text{CIS}} = (\mathbf{X}_{\text{new}}^{\text{CIS}} - \hat{\mu}^{\text{CIS}})^T \hat{\Omega}^{\text{CIS}} (\hat{\mu}_1^{\text{CIS}} - \hat{\mu}_2^{\text{CIS}})$ , then

## Proposition

*When classifying  $\mathbf{X}_{\text{new}}$  based on features selected from the screening step, we have*

$$R_{\text{CIS\_PSC}} - R_{\text{OPT}} \rightarrow_P 0$$



- A **discovery boundary**  $d(\beta)$  divides the  $\beta - r$  plane:
  - (a) If  $r > d(\beta)$ , it is possible to find a decision rule  $\psi$  such that at least some features in  $S_0$ , if not all, can be discovered by  $\psi(\tau)$
  - (b) If  $r < d(\beta)$ , it is impossible to find a decision rule  $\psi$  to find any informative features in  $S_0$
- A **classification boundary**  $c(\beta)$  divides the  $\beta - r$  plane:
  - (a) If  $r > c(\beta)$ , it is possible to find a decision rule  $\psi$  such that the two classes can be reliably separated
  - (b) If  $r < c(\beta)$ , it is impossible to find a decision rule  $\psi$  to reliably separate the two classes

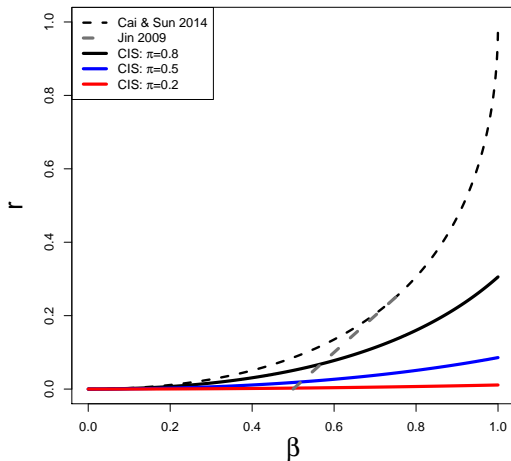
## Proposition (Discovery Boundary)

Let  $\pi = \min \left\{ \frac{\gamma - \log_p g(\alpha_{\min})}{\beta}, 1 \right\}$ , where  $\alpha_{\min} = \min_{(j,j') \in \mathcal{E}} |\Sigma_{jj'}|$  and  $g$  is an increasing function. Then the discovery boundary of the CIS classification rule  $\mathbf{1}(\hat{\psi}^{\text{CIS}}(\tau, \alpha, \nu) \geq 0)$  is always below that of the independence rule  $\mathbf{1}(\hat{\psi}^{\text{ind}}(\tau) \geq 0)$

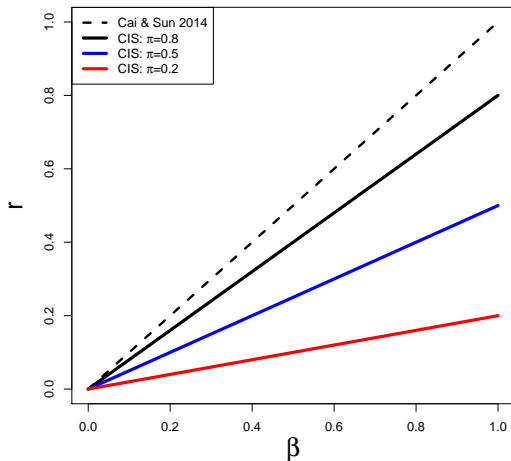
- $\pi\beta$  can be then interpreted as the effective sparsity of a marginal feature after taking into consideration its correlated MUJI features

## Proposition (Classification Boundary)

The classification boundary of the CIS classification rule  $\mathbf{1}(\hat{\psi}^{\text{CIS}}(\tau, \alpha, \nu) \geq 0)$  is also always below that of the independence rule  $\mathbf{1}(\hat{\psi}^{\text{ind}}(\tau) \geq 0)$



# Classification Boundaries



- 1 Background
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### 3-Class Classification

- 10,000 variables and 100 patients in each class
- The first 20 variables have mean structures between classes as in the table below
- The remaining 9,980 variables are independently and identically distributed from  $N(0, 1)$

**Mean of Variables**

Variables	Class 1	Class 2	Class 3
$X_1 - X_4, X_{11} - X_{14}$	0	0	-2.5
$X_5, X_{15}$	-0.5	2	-2.5
$X_6 - X_{10}, X_{16} - X_{20}$	1.5	-1.5	-1.5



: Marginal informative feature



: MUJI feature

# Simulation Results

When marginally informative and MUJI signals are correlated with a compound symmetry (CS) structure with a correlation = 0.5

Method	class pair	FP#	FN#	se	sp	MMS	ER
$\alpha = 0.2$							
CIS	1-2	1.1	0.1	0.994	0.999	24	1.4
	2-3	9.3	0	1	0.999	19	3.1
MS	1-2	16.4	9.1	0.54	0.998	8685	3.1
	2-3	3.0	0	1	0.999	10	3.4
$\alpha = 0.5$							
CIS	1-2	11.8	2.5	0.88	0.999	48	2.7
	2-3	5.6	0	1	0.999	10	2.9
MS	1-2	18.0	9.0	0.55	0.998	8634	3.0
	2-3	3.5	0	1	0.999	10	3.4
$\alpha = 0.9$							
CIS	1-2	12.2	7.8	0.61	0.999	8698	1.6
	2-3	2.7	0	1	0.999	10	3.4
MS	1-2	17.3	9.1	0.54	0.998	8876	3.1
	2-3	2.7	0	1	0.999	10	3.4

- 1 Background
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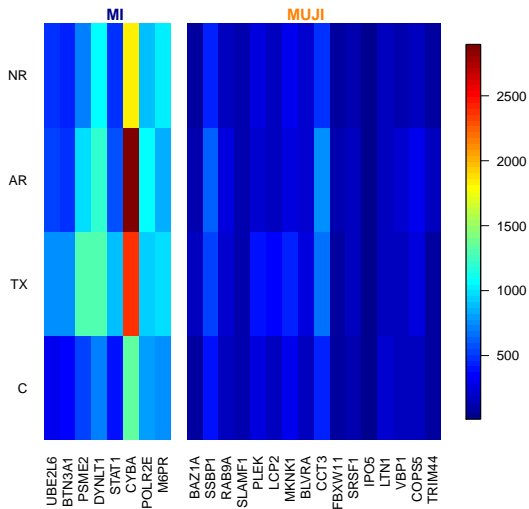


# Genomic Classification of Kidney Post-transplant Rejection Types

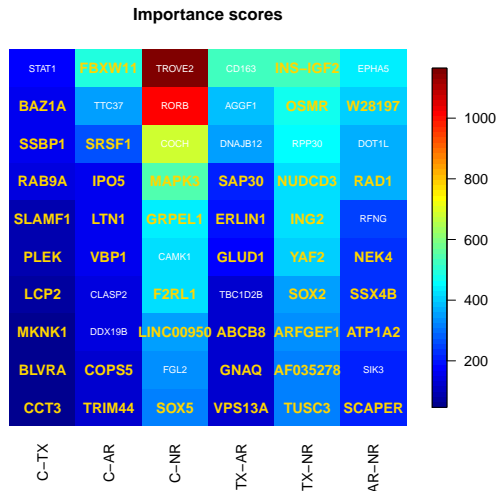
- 62 kidney tissue samples
- Four rejection types:
  - C – 17 normal donor kidneys
  - TX – 19 well-functioning transplants without rejections
  - AR – 13 acute rejection
  - ADNR – 13 acute dysfunction without rejection
- 12,625 gene expressions from patient kidney biopsies before transplant

- The same data were analyzed in [Xu et al. 2014](#)
- [Xu et al. 2014](#) method can not handle ultrahigh-dimensional data. They used an ad-hoc way to pre-select a subset of genes, and only involved 200 genes with the largest variations
- This could miss many MUJI genes
- We applied the CIS procedure for the purposes of MUJI gene identification and post-transplantation rejection type classification

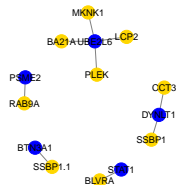
Mean expression scores for top marginally informative and MUJI genes



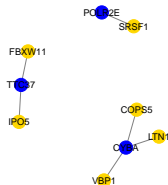
# Results(Cont.)



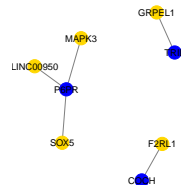
# Results (Cont.)



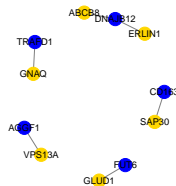
**C-TX**



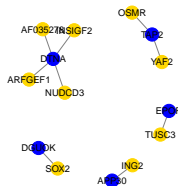
**C-AR**



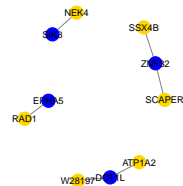
**C-NR**



**TX-AR**

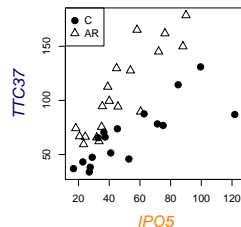
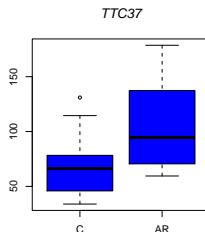
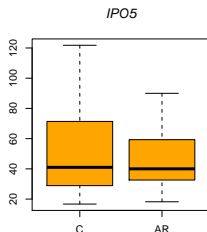


**TX-NR**



**AR-NR**

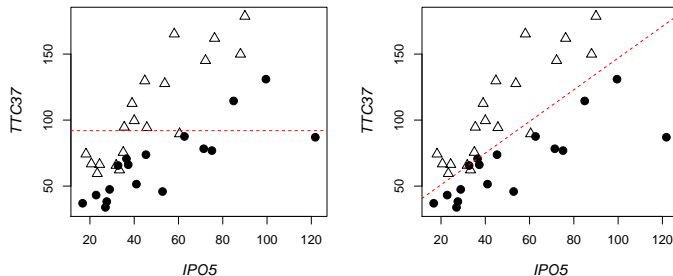
## gene *IPO5*



gene $j$	$\hat{\delta}_j^*$ of gene	rank of gene	correlated gene	cor- relation	rank of $IS_j$
<i>IPO5</i>	0.20	10279th	<i>TTC37</i>	0.7	4th

\*  $\hat{\delta}_j$  = standardized marginal mean difference of gene  $j$ .

## Results (Cont.)



- Left panel: based only on *TTC37*, the minimal misclassification number is **10**
- Right panel: based on *TTC37* and its correlated MUJI gene *IPO5*, the minimal misclassification number is **2**

- Use the leave-one-out procedure to predict each individual's class
  - Leave  $i$ th subject out at a time and use the rest as the training set
  - Apply the classification rule  $\mathbf{1}(\hat{\delta}^{\text{CIS}}(\mathbf{X}_i; \rho, \tau, m) \geq 0)$  to predict the label  $Y_i$  between each pair of classes
  - Take a majority voting on the final predicted class for subject  $i$
- In total 6 misclassifications out of 62 subjects



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- Leverages inter-feature correlations but circumvent estimation of the full precision matrix
- Works well with sparse and weak signals. Can effectively select both marginally informative and MUJI signals
- With nice theoretical properties, including optimal post-screening misclassification rates and improved discovery and classification boundaries
- Can we go beyond LDA?



# Thank You!

## Definition

A random variable  $X$  is sub-Gaussian with variance proxy  $\sigma^2$  if its moment generating function satisfies

$$E[\exp(sX)] \leq \exp\left(\frac{\sigma^2 s^2}{2}\right)$$

for any  $s \in \mathbb{R}$ .

A random vector  $\mathbf{X} = (X_1, \dots, X_p)$  is said to be sub-Gaussian with variance proxy  $\sigma^2$  if for any  $\mathbf{a} \in \mathbb{R}^p$ ,

$$E[\exp(\mathbf{a}'\mathbf{X})] \leq \exp\left(\frac{\sigma^2 \|\mathbf{a}\|_2^2}{2}\right).$$

- $\mathcal{G}(\mathcal{V}, \mathcal{E}; \mathbf{A})$  the graph induced by a  $p \times p$  symmetric matrix  $\mathbf{A}$ , where the node set  $\mathcal{V} = \{1, \dots, p\}$  corresponds to the row indices of  $\mathbf{A}$  and the edge set  $\mathcal{E}$  contains all the edges. An edge between nodes  $j$  and  $j'$  exists if and only if  $A_{jj'} \neq 0$
- Denote  $\mathcal{G}(\mathcal{V}_l, \mathcal{E}_l, \mathbf{A}_l)$  the connected components of  $\mathcal{G}(\mathcal{V}, \mathcal{E}; \mathbf{A})$

## Additional Simulation Results

When marginally informative and MUJI signals are correlated with a AR(1) structure and a correlation coefficient = 0.5

Method	class pair	FP#	FN#	se	sp	MMS	ER
$\alpha = 0.2$							
CIS	1-2	0.9	<b>0.05</b>	0.997	0.999	<b>31</b>	0.2
	2-3	9.9	0	1	0.999	17	0.9
MS	1-2	19.9	<b>9.6</b>	0.52	0.998	<b>8774</b>	2.1
	2-3	4.5	0	1	0.999	10	1.2
$\alpha = 0.5$							
CIS	1-2	9.1	<b>5.3</b>	0.74	0.999	<b>127</b>	0.6
	2-3	10.9	0	1	0.999	20	1.4
MS	1-2	11.2	<b>11.0</b>	0.45	0.999	<b>9354</b>	3.3
	2-3	4.6	0	1	0.999	10	1.4
$\alpha = 0.9$							
CIS	1-2	1.3	<b>9.1</b>	0.54	0.999	<b>8247</b>	0.4
	2-3	0.2	0.008	0.999	0.999	10	0.9
MS	1-2	1.9	<b>11.7</b>	0.41	0.999	<b>7470</b>	5.6
	2-3	0.2	0.02	0.998	0.999	10	1.0

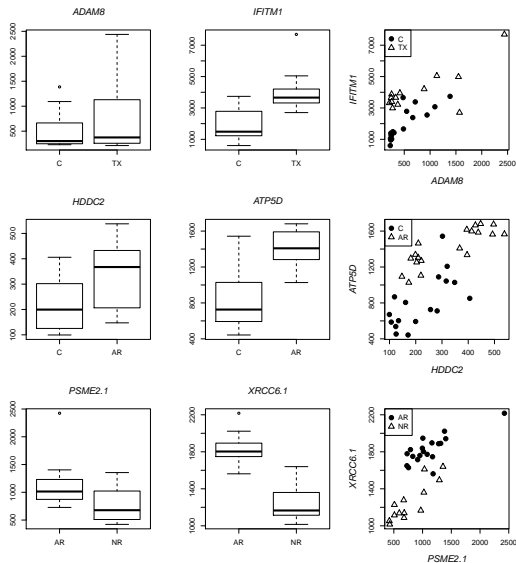
Comparisons of computing time (in minutes)

Methods	CIS-classification	ROAD	LPD	CED
$p = 200$	0.4	1.6	2.5	11.5
$p = 1000$	3.0	19.0	36.5	NA
$p = 10000$	4.5	NA	NA	NA
$p = 50000$	12.0	NA	NA	NA

\* Computational time longer than 5 hours is reported as NA.



# More Examples of MUJI Genes Detected



# More Examples of MUJI Genes Detected

Marginal and joint effects of MUJI genes

MUJI gene	$\hat{\delta}_j^*$ of MUJI gene	rank of MUJI gene	corelated MI** gene	$\hat{\delta}_j$ of MI gene	rank of MI gene	$\hat{\rho}$	rank of $IS_j$
<i>ADAM8</i>	-0.48	5071	<i>IFITM1</i>	-1.35	22	0.70	18
<i>HDDC2</i>	-0.82	2497	<i>ATP5D</i>	-1.53	51	0.71	77
<i>PSME2.1</i>	0.84	3962	<i>XRCC6.1</i>	1.71	22	0.74	22

\*  $\hat{\delta}_j$  = standardized marginal mean difference of gene j. \*\* MI= marginally informative