Fluctuations in a general preferential attachment model

Hanna Döring

Symposium in Memory of Charles Stein 1920 - 2016

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 n vertices, independently for each edge: Ber(p).
- The random portion of vertices with degree k converges with $p = \frac{\lambda}{n}$ to

$$\frac{1}{n}\sum_{i=1}^{n}\mathbb{1}_{\{\deg_n(i)=k\}}\xrightarrow{n\to\infty}p_k=e^{-\lambda}\frac{\lambda^k}{k!}$$

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 $\rightarrow\,$ BARABÁSI AND ALBERT (1999) proposed a different dynamic model

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The indegree distribution of a uniformly chosen vertex converges to a limiting distribution

$$\mathbb{P}\big(deg^{-}(U_n)=k\big) \xrightarrow{n \to \infty} \mu_k = \frac{1}{1+f(k)} \prod_{i=0}^{k-1} \frac{f(i)}{1+f(i)}, \quad k \in \mathbb{N}_0$$

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satisfying a power law for $f(k) = \gamma k + \beta$, $\gamma, \beta \in [0, 1)$ or a stretched exponential for $f(k) = \gamma k^{\alpha}$, $\alpha \in (0, 1), \gamma > 0$

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- G_n consists of n vertices, no loops, no multiple edges
 The new vertex connects to each of the former vertices j by at most one edge with probability

$$\mathbb{P}(n+1 \text{ connects to } j | \mathcal{G}_n) = \frac{f(\deg_n^-(j))}{n},$$

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where deg_n⁻(j) is the indegree of vertex j at time n and $f : \mathbb{N}_0 \to (0, \infty)$ is monotonically increasing such that $f(n) \leq n+1$.

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Preferential Attachment Model: Examples

- fixed outdegree exactly one edge each step:
 - $d_0 = 1$, add one vertex and set $f(k) = \frac{k+1+\delta}{n(2+\delta)}$ for a fixed parameter $\delta > -1$.

close to Bollobás, Riordan, Spencer and Tusnády (2001); Hofstad (2017)

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 $d_0 = 0, n+1$ has an edge with j < n+1 independently with prob. $\frac{f(\deg_n^-(j))}{n}$

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• spatial attachment:

on a hypercube; add an edge with probability p if it lies in the *sphere* of influence of an older vertex which itself depends on the degree AIELLO, BONATO, COOPER, JANSSEN AND PRALAT (2008)

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Theorem 1 (Betken, D., Ortgiese, 2019)

Let W_n denote the indegree of a uniformly chosen vertex at time n in a preferential attachment model satisfying our assumptions. Suppose further that there exists $k_* \in \mathbb{N}_0$ such that f(k) > k for all $k < k_*$ and $f(k) \le k$ for all $k \ge k_*$.

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Let W_n denote the indegree of a uniformly chosen vertex at time n in a preferential attachment model satisfying our assumptions. Suppose further that there exists $k_* \in \mathbb{N}_0$ such that f(k) > k for all $k < k_*$ and $f(k) \le k$ for all $k \ge k_*$.

Then, there exists a constant C > 0 such that for all $n \ge 2$

$$d_{\mathrm{TV}}(W_n, W) \leq C \ \frac{\log(n)}{n},$$

where
$$W \sim \mu = (\mu_k)_k = \left(\frac{1}{1+f(k)} \prod_{i=0}^{k-1} \frac{f(i)}{1+f(i)}\right)_k$$
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$$\mathbb{P}(N_Y \ge k) = \mathbb{P}(Y \ge S_k) = \mathbb{E}\big[\mathbb{E}[\mathbb{1}_{\{Y \ge S_k\}} | S_k]\big] = \mathbb{E}\big[e^{-S_k}\big]$$

• DEREICH, MÖRTERS (2009): limiting distribution μ

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$$\mathbb{P}(N_Y = k) = \prod_{i=0}^{k-1} \frac{f(i)}{1 + f(i)} - \prod_{i=0}^k \frac{f(i)}{1 + f(i)} = \mu_k$$

Fluctuations in a general PA model

• for all
$$g : \mathbb{N}_0 \to \mathbb{R}$$

$$\mathbb{E}[g(N_Y)] = \int_0^\infty \mathbb{E}[g(N_s)]e^{-s}ds$$

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= $g(0) + \mathbb{E}[f(N_Y)(g(N_Y+1) - g(N_Y))].$

• Stein operator: Choose

$$\begin{aligned} \mathcal{A}g(k) &:= f(k) \big(g(k+1) - g(k) \big) + g(0) - g(k) \\ &= f(k) \Delta g(k) + g(0) - g(k) \end{aligned}$$

for $\Delta g(k) = g(k+1) - g(k)$.

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Stein Equation for PA

With

$$\mathcal{A}g(k) = f(k)\Delta g(k) + g(0) - g(k)$$

and

$$\mu = (\mu_k) = \left(\frac{1}{1+f(k)}\prod_{i=0}^{k-1}\frac{f(i)}{1+f(i)}\right)_k$$

we have

$$W \sim \mu \quad \Leftrightarrow \quad \mathbb{E}[\mathcal{A}g(W)] = 0$$

for all $g: \mathbb{N}_0 \to \mathbb{R}$ such that $\mathbb{E}[g(W)] < \infty$.

Solution to Stein's Equation for PA

Let (Z_t) denote a Markov chain with generator A solution of Ag_A = h − μ(h) for any h = 1_A, with A ⊂ N₀ is given by

$$g_A(k) = -\int_0^\infty \left(\mathbb{E}_k h(Z_t) - \int h d\mu\right) dt$$

if it exists.

BARBOUR (1988), GÖTZE (1991)

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Solution to Stein's Equation for PA

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if it exists.

BARBOUR (1988), GÖTZE (1991)

- here: The solution can be checked using Kolmogorov's backward equation for Markov chains and a small coupling argument
- smoothness estimate:

$$f(k)\Delta g(k)\leq 1$$
 for all $k\in\mathbb{N}_0$

following techniques in BROWN, XIA (2001)

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• dynamic way of generating a uniform random variable on $\{1, \ldots, n\}$ cf. FORD (2009)

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- dynamic way of generating a uniform random variable on $\{1, \ldots, n\}$ cf. FORD (2009)
- Let J_n be a Markov chain with $J_1 = 1$ and such that

$$\mathbb{P}(J_{n+1} = J_n | J_n) = \frac{n}{n+1}$$
 and $\mathbb{P}(J_{n+1} = n+1 | J_n) = \frac{1}{n+1}$

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• For any $n, J_n \sim U(\{1, \dots, n\})$

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• $W_n := \deg^-(U_n) \stackrel{d}{=} X_n := \deg^-(J_n)$

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For any $n, J_n \sim U(\{1, \dots, n\})$

$$W_n := \deg^-(U_n) \stackrel{d}{=} X_n := \deg^-(J_n)$$

• (X_n) is a time-inhomogeneous, discrete Markov chain, with generator

$$\frac{f(k)}{n+1}\Delta g(k) + \frac{1}{n+1}(g(0) - g(k))$$

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$$W_n \stackrel{d}{=} X_n := \deg^-(J_n)$$

• transition probabilities given for any $i \ge 1$ as

$$\mathbb{P}(X_{n+1} = j | X_n = i) = \begin{cases} \frac{f(i)}{n+1} & \text{if } j = i+1, \\ & & \\$$

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and

$$\mathbb{P}(X_{n+1} = j | X_n = 0) = \begin{cases} \frac{f(0)}{n+1} & \text{if } j = 1, \end{cases}$$

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$$W_n \stackrel{d}{=} X_n := \deg^-(J_n)$$

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and

$$\mathbb{P}(X_{n+1} = j \mid X_n = 0) = \begin{cases} \frac{f(0)}{n+1} & \text{if } j = 1, \\ 1 - \frac{f(0)}{n+1} & \text{if } j = 0. \end{cases}$$

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• aim: Bound

$$\begin{split} |\mathbb{E}[\mathcal{A}g_{\mathcal{A}}(\mathcal{W}_{n+1})]| &= |\mathbb{E}\mathcal{A}g_{\mathcal{A}}(X_{n+1})]| \\ &= |\mathbb{E}[f(X_{n+1})\Delta g_{\mathcal{A}}(X_{n+1}) + g_{\mathcal{A}}(0) - g_{\mathcal{A}}(X_{n+1})]| \end{split}$$

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• Let $h : \mathbb{N}_0 \to \mathbb{R}$ be such that h(0) = 0, then $\mathbb{E} \left[h(X_{n+1}) \right] = \mathbb{E} \left[\mathbb{E} \left[h(X_{n+1}) \mid X_n \right] \right]$

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$$\mathbb{E}[h(X_{n+1})] = \mathbb{E}[\mathbb{E}[h(X_{n+1}) | X_n]]$$

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= $\frac{n}{n+1} \mathbb{E}[h(X_n)] + \frac{1}{n+1} \mathbb{E}[f(X_n)\Delta h(X_n)].$

• iteration and $X_1 = 0$ yields

$$\mathbb{E}\left[h(X_{n+1})\right] = \frac{1}{n+1} \sum_{\ell=1}^{n} \mathbb{E}[f(X_{\ell})\Delta h(X_{\ell})].$$

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for $h(k) = \mathcal{A}g_{\mathcal{A}}(k) - v_{\mathcal{A}}(0) = v_{\mathcal{A}}(k) + g_{\mathcal{A}}(0) - g_{\mathcal{A}}(k) - v_{\mathcal{A}}(0)$:

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$$\mathbb{E}[h(X_{n+1})] = \frac{1}{n+1} \sum_{\ell=1}^{n} \mathbb{E}[f(X_{\ell})\Delta h(X_{\ell})]$$

for $h(k) = \mathcal{A}g_{A}(k) - v_{A}(0) = v_{A}(k) + g_{A}(0) - g_{A}(k) - v_{A}(0)$:
 $\mathbb{E}[\mathcal{A}g_{A}(X_{n+1})] = \mathbb{E}[h(X_{n+1})] + v_{A}(0)$
 $= \frac{1}{n+1} \sum_{\ell=1}^{n} \left(\mathbb{E}[f(X_{\ell})\Delta v_{A}(X_{\ell})] - \mathbb{E}[f(X_{\ell})\Delta g_{A}(X_{\ell})] \right) + v_{A}(0)$
 $= \frac{1}{n+1} \sum_{\ell=1}^{n} \sum_{k=0}^{\ell-1} f(k)\Delta v_{A}(k)\mathbb{P}(X_{\ell} = k)$
 $- \frac{1}{n+1} \sum_{\ell=1}^{n} \sum_{k=0}^{\ell-1} (v_{A}(k) - v_{A}(0))\mathbb{P}(X_{\ell} = k) + \frac{v_{A}(0)}{n+1}.$

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• For the second sum, we write

$$\sum_{k=0}^{\ell-1} (v_A(k) - v_A(0)) \mathbb{P}(X_\ell = k) = \sum_{k=0}^{\ell-1} \sum_{i=0}^{k-1} \Delta v_A(i) \mathbb{P}(X_\ell = k) = \sum_{i=0}^{\ell-1} \Delta v_A(i) \sum_{k=i+1}^{\ell-1} \mathbb{P}(X_\ell = k) = \sum_{i=0}^{\ell-1} \Delta v_A(i) \mathbb{P}(X_\ell \ge i+1).$$

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• For the second sum, we write

$$\sum_{k=0}^{\ell-1} (v_A(k) - v_A(0)) \mathbb{P}(X_\ell = k) = \sum_{k=0}^{\ell-1} \sum_{i=0}^{k-1} \Delta v_A(i) \mathbb{P}(X_\ell = k) = \sum_{i=0}^{\ell-1} \Delta v_A(i) \sum_{k=i+1}^{\ell-1} \mathbb{P}(X_\ell = k) = \sum_{i=0}^{\ell-1} \Delta v_A(i) \mathbb{P}(X_\ell \ge i+1).$$

• Therefore,

$$\mathbb{E}[\mathcal{A}g_{\mathcal{A}}(X_{n+1})] = \frac{1}{n+1} \sum_{\ell=1}^{n} \sum_{k=0}^{\ell-1} f(k) \Delta v_{\mathcal{A}}(k) \mathbb{P}(X_{\ell} = k) \\ - \frac{1}{n+1} \sum_{\ell=1}^{n} \sum_{k=0}^{\ell-1} (v_{\mathcal{A}}(k) - v_{\mathcal{A}}(0)) \mathbb{P}(X_{\ell} = k) + \frac{v_{\mathcal{A}}(0)}{n+1}$$

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$$= \sum_{i=0}^{\ell-1} \Delta v_A(i) \sum_{k=i+1}^{\ell-1} \mathbb{P}(X_\ell = k) = \sum_{i=0}^{\ell-1} \Delta v_A(i) \mathbb{P}(X_\ell \ge i+1).$$

• Therefore,

$$\mathbb{E}[\mathcal{A}g_{\mathcal{A}}(X_{n+1})] = \frac{1}{n+1} \sum_{\ell=1}^{n} \sum_{k=0}^{\ell-1} f(k) \Delta v_{\mathcal{A}}(k) \mathbb{P}(X_{\ell} = k) \\ - \frac{1}{n+1} \sum_{\ell=1}^{n} \sum_{k=0}^{\ell-1} \Delta v_{\mathcal{A}}(k) \mathbb{P}(X_{\ell} \ge k+1) + \frac{v_{\mathcal{A}}(0)}{n+1}$$

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• Thus, we used the Markov structure to show

$$\mathbb{E}\left[\mathcal{A}g_{\mathcal{A}}(X_{n+1})\right] = \frac{1}{n+1} \Big(\sum_{\ell=1}^{n} \sum_{k=0}^{\ell-1} \Delta v_{\mathcal{A}}(k)h(k,\ell) + v_{\mathcal{A}}(0)\Big),$$

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• By discrete integration by parts formula and $h(\ell, \ell) = 0$:

$$\sum_{k=0}^{\ell-1} \Delta v_{\mathcal{A}}(k) h(k,\ell) = -v_{\mathcal{A}}(0) h(0,\ell) - \sum_{k=0}^{\ell-1} v_{\mathcal{A}}(k+1) \Delta_k h(k,\ell).$$

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$$\begin{split} & \left| \mathbb{E}[\mathcal{A}g_{A}(X_{n+1})] \right| \\ & \leq \frac{|v_{A}(0)|}{n+1} + \left| \frac{1}{n+1} \sum_{\ell=1}^{n} \left(\left(\sum_{k=0}^{\ell-1} v_{A}(k+1) \Delta_{k} h(k,\ell) \right) + v_{A}(0) h(0,\ell) \right) \right| \\ & \leq \frac{|v_{A}(0)|}{n+1} + \frac{2}{n+1} \sup_{k} |v_{A}(k)| \sum_{\ell=1}^{n} \sup_{k \leq \ell-1} h(k,\ell). \end{split}$$

And finally

$$d_{\mathrm{TV}}(W_n, W) = \sup_{A \subset \mathbb{N}_0} \left| \mathbb{E}[\mathcal{A}g_A(X_{n+1})] \right|$$

$$\leq \frac{1}{n+1} + \frac{2}{n+1} \sum_{\ell=1}^n \sup_{k \leq \ell-1} h(k, \ell)$$

$$\leq \begin{cases} C \frac{\log(n)}{n} & \text{for slowly growing } f \\ C \frac{1}{n^{1-\gamma}} & \text{for } f(k) = k + \gamma. \end{cases}$$

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- Known: $\lambda_n \to \lambda := \mathbb{E}[f(W)]$, where $W \sim \mu$. DEREICH, MÖRTERS (2009)
- If $f(k) = \gamma k + \beta$ for $\gamma \in (0,1), \beta \in [0,1]$, then

$$|\lambda_n - \mathbb{E}[f(W)]| \leq n^{-1+\gamma}$$

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we have

$$\lambda_n := \mathbb{E}[D_n] = \mathbb{E}\Big[\frac{1}{n-1}\sum_{i=1}^{n-1}f(\deg_{n-1}^-(i))\Big] = \mathbb{E}\left[f(W_{n-1})\right].$$

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• Stein's equation

 $\mathbb{E}[\lambda g(Z+1) - Zg(Z)] = 0 \text{ for all bounded } g: \mathbb{N} \to \mathbb{R} \iff Z \sim \textit{Po}(\lambda)$

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$$d_{\mathrm{TV}}(D_n, \mathsf{Po}(\lambda_n)) \leq \frac{1 - e^{-\lambda_n}}{\lambda_n} \sum_{i=1}^{n-1} p_{i,n}^2 \leq \min\left\{1, \frac{1}{\lambda_n}\right\} \sum_{i=1}^{n-1} p_{i,n}^2.$$

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here:

$$\sum_{i=1}^{n-1} p_{i,n}^2 = \frac{1}{(n-1)^2} \sum_{i=1}^{n-1} \mathbb{E}[f(\deg_{n-1}^{-}(i))]^2 \stackrel{(*)}{\leq} \frac{1}{(n-1)^2} \sum_{i=1}^{n-1} \left(\frac{n}{i}\right)^{2\gamma},$$

(*): see DEREICH, MÖRTERS (2013).

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